

The urban gender premium in intergenerational mobility

Evidence from South Asia

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19th Winter School On Inequality and Social Welfare Theory
Canazei, Jan 7-9, 2026

Intro

Motivation

- Developing countries tend to exhibit lower levels of social mobility (Narayan and Van der Weide, 2018).
- Spatial variation: Urban areas tend to be sources of economic opportunity, shifting social norms (Fafchamps and Shilpi, 2005; Bussolo et al., 2022).
- This might be particularly true for women.

Question: Do cities contribute to upward mobility in developing countries? Why?

Empirical challenge:

- Lack of linked parent-child income data makes estimating IGM a challenge→ most work focused on developed countries (Chetty et al., 2014a,b).
- Parent-child linked data used for developing countries are educational data, which tend to be coarse, and no standardized measure of mobility was suitable for comparative analysis in these situations.
- Bottom Half Educational Mobility (μ_0^{50}) fills this gap.

Context: South Asia

- Rapid growth and urbanization: since 1990, urban population has grown by 380 million, likely underestimated (World Bank).
- Yet persistent socioeconomic and labor market inequities across multiple dimensions: gender, caste, religion (Munshi, 2019; Klasen and Pieters, 2015).
- Highest intergenerational rank persistence across the world (Narayan and Van der Weide, 2018).
- **But limited evidence on variation in mobility within the region**

This paper

We investigate the role of gender and urbanization in driving educational mobility, extending estimates from Asher et al. (2024) to all of South Asia.

Methods:

- Build harmonized microdata spanning 5.9 million South Asians across 7 countries.
- Match parent–child educational attainment from 39 nationally representative household surveys.
- Construct comparable location-specific mobility measures for boys and girls.

We identify the following empirical facts:

- Social mobility generally low across South Asia, though with heterogeneity.
- Urban mobility premium (UMP): social mobility higher in cities.
- Female urban mobility premium (FUMP): Women have higher UMP than men.

Methods

Data

Table 1: Datasets Harmonized for Analysis

| Country | Survey Name | Years of Data | N: HH | Co-resident |
|-------------|---|--------------------------------------|--------|-----------------|
| Afghanistan | Afghanistan Living Conditions Survey | 2013, 16 | 40607 | Yes |
| Afghanistan | National Risk & Vulnerability Assessment | 2008, 12 | 41404 | Yes |
| Afghanistan | Income, Expenditure & Labor Force Survey | 2019 | 18344 | Yes |
| Sri Lanka | Household Income & Expenditure Survey | 1991, 95, 2002, 06, 09, 12 | 95719 | Yes |
| Nepal | Nepal Living Standard Survey | 1995, 2003, 11 | 7316 | Yes (exc. 2011) |
| India | India Human Development Survey | 2012 | 38387 | No |
| Bangladesh | Household Income & Expenditure Survey | 2000, 05, 10, 16 | 75828 | Yes |
| Bhutan | Bhutan Living Standards Survey | 2003, 07, 12, 17 | 34415 | Yes (exc. 2003) |
| Pakistan | Pakistan Integrated Household Survey | 1991 | 4791 | No |
| Pakistan | Pakistan Social & Living standards Measurement Survey | 2006, 08, 10, 12, 14, 19 | 283426 | Yes |
| Pakistan | Household Income & Expenditure Survey | 2001, 04, 05, 07, 10, 11, 13, 15, 18 | 129906 | Yes |

Methods: sample selection

Requirement: Matched parent-child education pairs

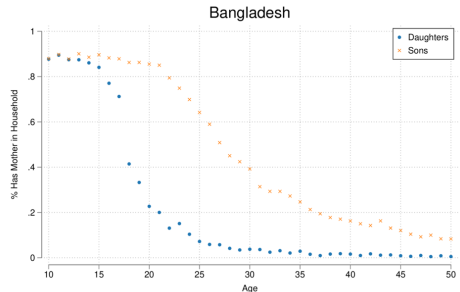
- In many datasets, we only observe parent-child pairs when they live together
- This results in bias if upwardly mobile children are more likely to stay or to exit

Our solution \Rightarrow Target ages when:

- Most children still at home
- Education attainment is stable:
i.e. High-school for boys,
Middle-school for girls

This solution works well for our countries

Figure 1: Co-residence age: Bangladesh



Co-residence Women

Methods: estimating bottom-half mobility

Define bottom-half mobility (BHM) as $\mu_0^{50} = E[\text{rank}_{child} | \text{rank}_{father} \leq 50]$

- same focus and same range of the p25 measure used by Chetty et al
 $p25 = E[\text{rank}_{child} | \text{rank}_{father} = 25]$
- Random mobility, $\mu_0^{50} = 50$; no mobility, $\mu_0^{50} = 25$

Issue: partial identification of μ_0^{50}

- Since education distribution is discrete, one attainment category might cover the 50th percentile \rightarrow *bounds* on μ_0^{50}

Estimation procedure:

1. Standardize educational attainment into 6-group categorical variable.
2. Rank sons/daughters and fathers within country-birth cohort.
3. Calculate bounds on μ_0^{50}

Key assumption! Monotonicity of rank-rank function.

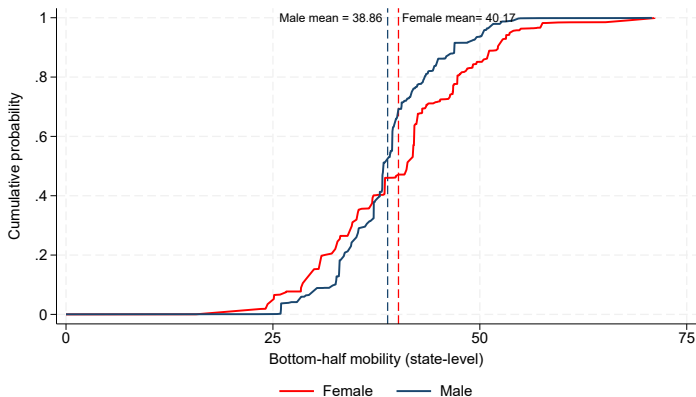
Bounds example

Education-consumption

Results

Bottom Half Mobility (BHM) in South Asia

Figure 2: Mobility distribution: state-level

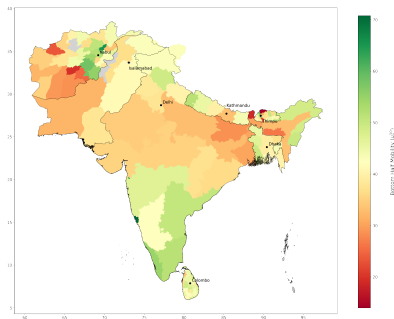


Cohort-level

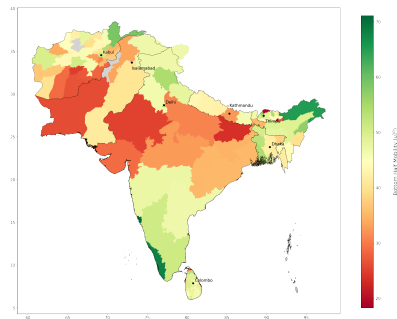
Rank-rank estimates

Country-level

Mobility map: state-level



(a) Sons



(b) Daughters

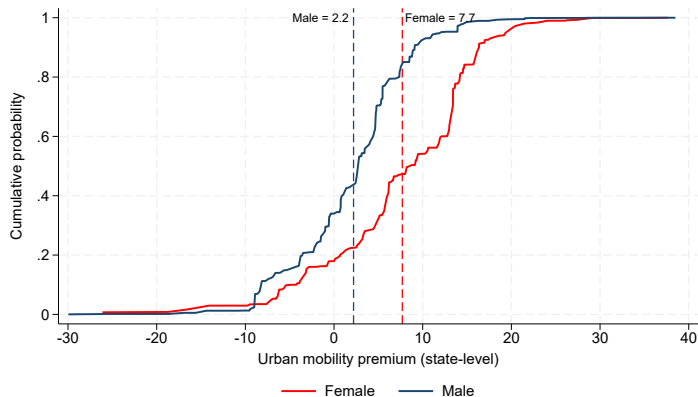
BHM & Urb Mobility Premium (UMP) — Country-level midpoint estimates

Table 2: Bottom half mobility

| | Female | | | Male | | |
|-------------|--------|-------|-------|-------|-------|-------|
| | Urban | Rural | UMP | Urban | Rural | UMP |
| Afghanistan | 56.75 | 38.73 | 18.02 | 51.44 | 38.78 | 12.66 |
| Bangladesh | 42.97 | 45.38 | -2.41 | 40.79 | 41.58 | -0.80 |
| Bhutan | 57.50 | 45.31 | 12.18 | 58.99 | 44.34 | 14.65 |
| India | 47.52 | 34.91 | 12.61 | 40.40 | 36.98 | 3.42 |
| Nepal | 52.31 | 40.60 | 11.71 | 47.53 | 38.32 | 9.21 |
| Pakistan | 48.24 | 35.21 | 13.04 | 40.87 | 36.95 | 3.91 |
| Sri Lanka | 45.14 | 46.36 | -1.22 | 45.68 | 43.50 | 2.19 |

The UMP

Figure 4: Urban mobility premium: state-level



Note: Figure shows the cumulative distribution function of state-level urban bottom half mobility premia for men and women in South Asia. Sample is all cohorts since 1980. Distribution is weighted by state-level population. Vertical lines show average urban premia in bottom half mobility. Mobility estimates are taken at midpoints of bounds.

Regression-based estimates of UMP

Table 3: Rank-level mobility regression

| Mobility measure | Bottom-half | | | | Top-half | Level |
|------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| Sample | Male | Female | All | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Urban | 3.610*** (0.508) | 10.488*** (0.565) | 6.725*** (0.399) | 3.610*** (0.508) | 8.457*** (0.384) | 0.057*** (0.010) |
| Female | | | | -1.032** (0.446) | -1.209*** (0.361) | -0.136*** (0.009) |
| Urban × Female | | | | 6.877*** (0.715) | 1.649*** (0.476) | 0.099*** (0.014) |
| Constant | 38.121*** (0.294) | 37.089*** (0.345) | 37.639*** (0.229) | 38.121*** (0.294) | 54.292*** (0.283) | 0.441*** (0.006) |
| Observations | 129280 | 97552 | 226832 | 226832 | 315297 | 226832 |
| R ² | 0.003 | 0.029 | 0.012 | 0.015 | 0.034 | 0.023 |

Note: Standard errors in parentheses clustered at the household level. Estimates are from a regression of child education outcomes (rank or level) on indicator variables for urban residence and female identity. Sample and mobility outcomes are indicated in table header. "Bottom-half" and "level" columns restrict the sample to parents in the bottom 50% of educational attainment, while top-half restricts to parents in the top 50%. Sample is all cohorts since 1980 pooled across seven South Asian countries. All estimates adjusted with cross-country sampling weights. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

child rank = $\alpha + \beta_1 \times \text{urban} + \beta_2 \times \text{female} + \beta_3 \times \text{female} \times \text{urban}$, if parent rank ≤ 50

Drivers of FUMP

1) Returns to education

- Higher labor market returns to education for urban women may incentivize greater investment in education.
- In context with low FLFP, marriage market returns might also serve to incentivize educational investments.

Drivers of FUMP

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2) Social norms

- In places with less conservative norms, pre-existing social hierarchies are less likely to disproportionately constrain life outcomes for disadvantaged groups.

Drivers of FUMP

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- In places with less conservative norms, pre-existing social hierarchies are less likely to disproportionately constrain life outcomes for disadvantaged groups.

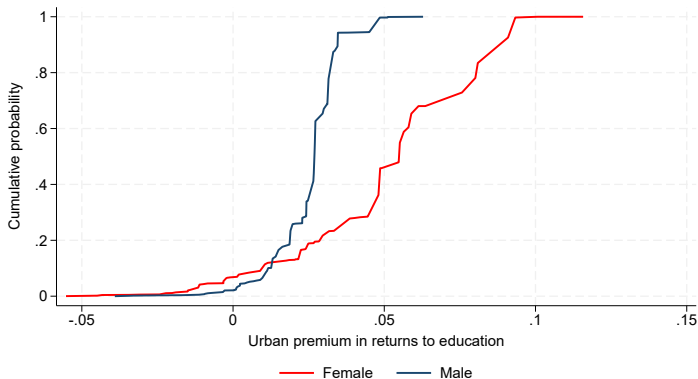
3) Income effects

- Higher incomes available in urban areas may increase educational investment, particularly for girls.

Note that in order to increase mobility, each of these explanations requires greater responses *by poor households*.

Returns to education: higher for urban women

Figure 5: Returns to education for men and women

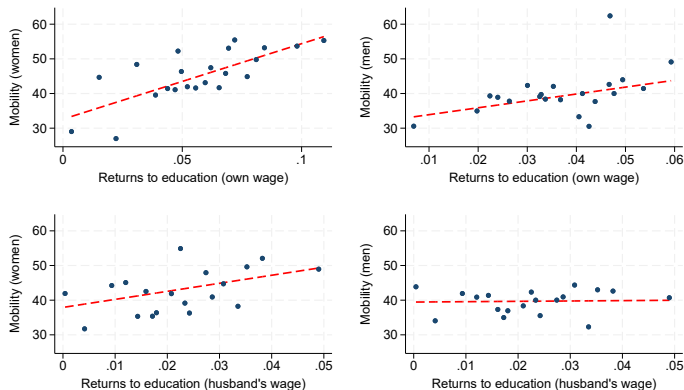


Cohorts since 1980, population weighted

Mincer regression

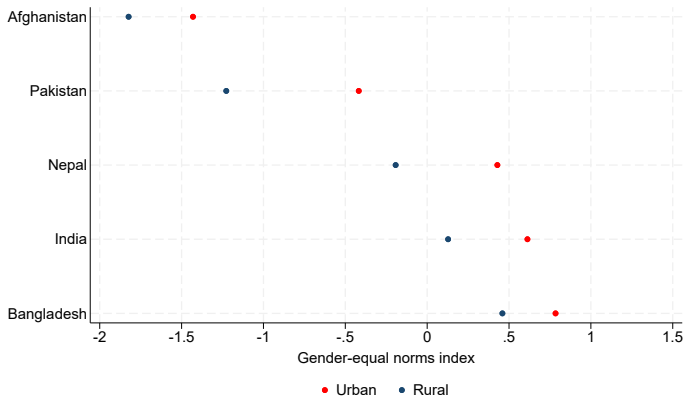
Returns to education: correlated with mobility

Figure 6: Mobility and returns to education



Gender norms: looser in urban areas

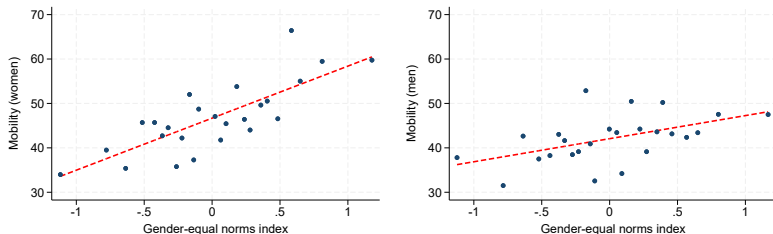
Figure 7: Urban and rural gender norms by country



Note: Figure shows average values of the standardized gender-equal norms index in rural and urban areas across countries. The gender-equal norms index is averaged across the following dimensions: son preference, reproductive choice, justification of domestic violence, and household decisionmaking. Each of these index components is measured as the share of female DHS respondents agreeing with a gender-equal norm. The resulting index, which varies between 0 and 1, is standardized across sub-national units. All estimates are sample-weighted.

Gender norms: correlated with female mobility

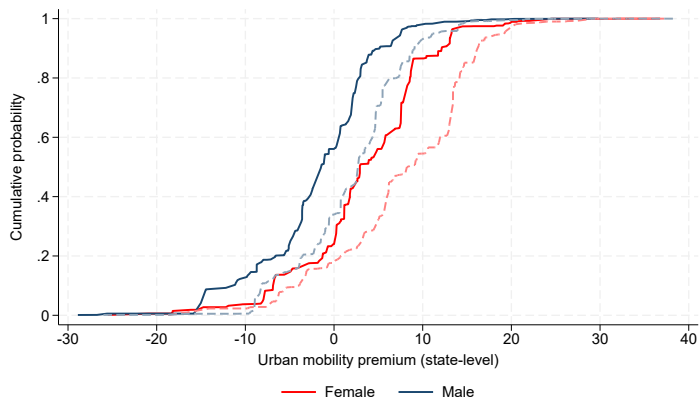
Figure 8: Mobility and gender norms



Regression test

Income effects

Figure 9: State-level UMP: Income-adjusted



Conclusion

- We introduce a comparable estimate of inter-generational educational mobility that can be constructed with standard household surveys.
- Mobility is low but spatially heterogeneous across South Asia.
- Women experience large mobility gains in cities.
- 65% of the FUMP is explained by returns to education and social norms.

UMP decomposition

Appendix

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- Chetty, Raj, Nathaniel Hendren, Patrick Kline, and Emmanuel Saez, "Where is the land of opportunity? The geography of intergenerational mobility in the United States," Quarterly Journal of Economics, 2014, 129 (4), 1553–1623.
- , —, —, —, —, and Nicholas Turner, "Is the United States Still a Land of Opportunity? Recent Trends in Intergenerational Mobility," American Economic Review: Papers & Proceedings, 2014, 104 (5), 141–147.
- Fafchamps, Marcel and Forhad Shilpi, "Cities and Specialisation: Evidence from South Asia," The Economic Journal, 2005, 115 (503), 477–504.
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Fair Progress? Economic Mobility Across Generations Around the World, World Bank Group, 2018.

Methods: bounds example

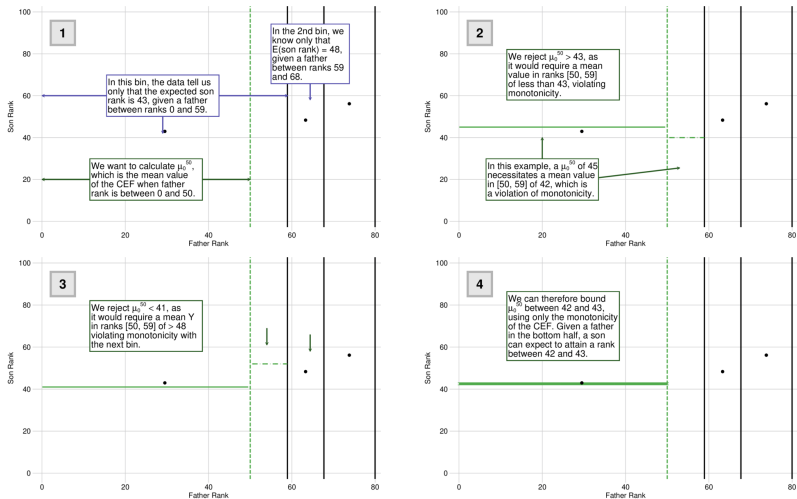
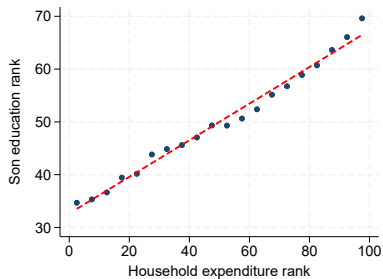
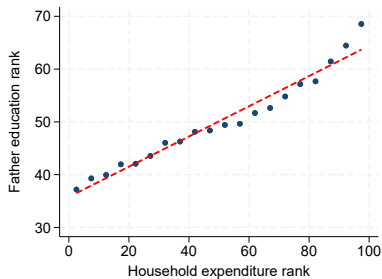


Figure A1: Calculating bounds on μ_0^{50}

Education vs. consumption

Figure A2: Consumption-education rank-rank correlations



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Coreidence: women

Figure A3: Coreidence and educational attainment: women

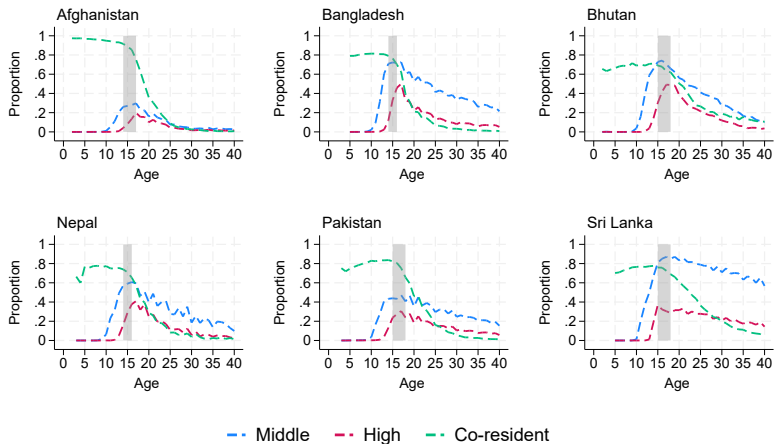


Figure A4: Coresidence and educational attainment: men

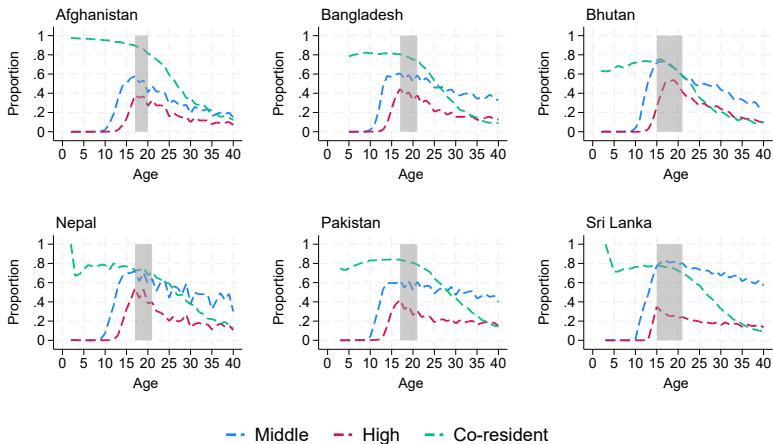


Table A1: Rank-rank mobility regression

| Dependent variable | Child's education rank | | | |
|--------------------------|------------------------|---------------------|---------------------|---------------------|
| Sample | Male | | Female | |
| | (1) | (2) | (3) | (4) |
| Father's education rank | 0.482*** (0.006) | 0.470*** (0.006) | 0.440*** (0.006) | 0.427*** (0.006) |
| Predicted BHM | 38.313 | 38.606 | 39.184 | 39.510 |
| Observations | 240617 | 240617 | 177950 | 177950 |
| R^2 | 0.228 | 0.255 | 0.210 | 0.289 |
| Cohort \times State FE | No | Yes | No | Yes |

Note: Standard errors in parentheses clustered at the household level. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Figure A5: Cohort-wise mobility estimates: country-level

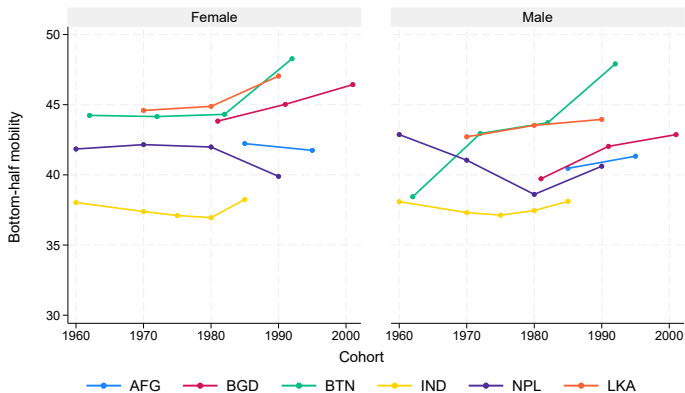
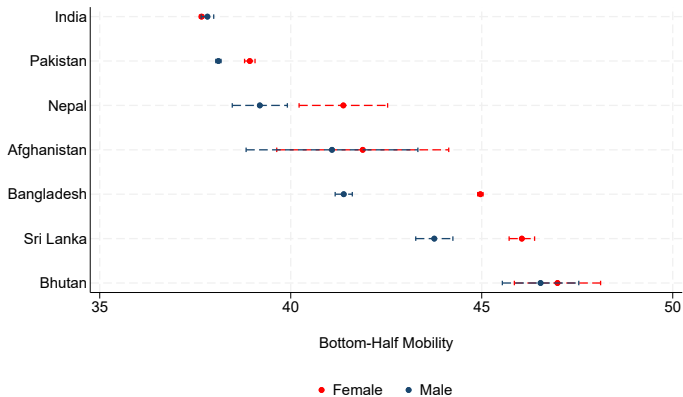


Figure A6: Country-level BHM



Rank-rank estimates of UMP

Table A2: Rank-rank mobility regression: UMP estimates

| Dependent variable Sample | Child's education rank | | | | |
|---|------------------------|---------------------|----------------------|----------------------|----------------------|
| | Male | Female | All | | |
| | (1) | (2) | (3) | (4) | (5) |
| Father's education rank | 0.426*** (0.009) | 0.449*** (0.009) | 0.449*** (0.009) | 0.435*** (0.009) | 0.373*** (0.020) |
| Urban | 12.529*** (0.812) | 1.275 (0.847) | 1.275 (0.847) | 0.489 (0.832) | 1.075 (0.873) |
| Father's education rank \times Urban | -0.066*** (0.012) | 0.046*** (0.013) | 0.046*** (0.013) | 0.050*** (0.013) | 0.041** (0.013) |
| Father's education rank \times Female | | | -0.023 (0.013) | -0.017 (0.012) | -0.002 (0.024) |
| Father's education rank \times Urban \times Female | | | -0.112*** (0.017) | -0.108*** (0.017) | -0.114*** (0.018) |
| Father's education rank \times Log per capita consumption | | | | | 0.009*** (0.002) |
| Father's education rank \times Log per capita consumption \times Female | | | | | -0.002 (0.003) |
| Female UMP | | | 9.603 | 8.712 | 9.715 |
| Male UMP | | | 2.423 | 1.735 | 2.096 |
| Observations | 177869 | 240617 | 418486 | 418486 | 397918 |
| R^2 | 0.235 | 0.233 | 0.234 | 0.269 | 0.237 |
| Cohort \times State FE | No | No | No | Yes | No |

Note: Standard errors in parentheses clustered at the household level. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Country-level midpoint estimates

Table A3: Bottom half mobility

| | Female | | | Male | | |
|-------------|--------|-------|-------|-------|-------|-------|
| | Urban | Rural | UMP | Urban | Rural | UMP |
| Afghanistan | 56.75 | 38.73 | 18.02 | 51.44 | 38.78 | 12.66 |
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| Nepal | 52.31 | 40.60 | 11.71 | 47.53 | 38.32 | 9.21 |
| Pakistan | 48.24 | 35.21 | 13.04 | 40.87 | 36.95 | 3.91 |
| Sri Lanka | 45.14 | 46.36 | -1.22 | 45.68 | 43.50 | 2.19 |

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Regression-based estimates of UMP

Table A4: Rank-level mobility regression

| Mobility measure | Bottom-half | | | | Top-half | Level |
|------------------|----------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| Sample | Boys | Girls | All | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Urban | 3.610*** (0.508) | 10.488*** (0.565) | 6.725*** (0.399) | 3.610*** (0.508) | 8.457*** (0.384) | 0.057*** (0.010) |
| Female | | | | -1.032** (0.446) | -1.209*** (0.361) | -0.136*** (0.009) |
| Urban × Female | | | | 6.877*** (0.715) | 1.649*** (0.476) | 0.099*** (0.014) |
| Constant | 38.121*** (0.294) | 37.089*** (0.345) | 37.639*** (0.229) | 38.121*** (0.294) | 54.292*** (0.283) | 0.441*** (0.006) |
| Observations | 129280 | 97552 | 226832 | 226832 | 315297 | 226832 |
| R^2 | 0.003 | 0.029 | 0.012 | 0.015 | 0.034 | 0.023 |

Note: Standard errors in parentheses clustered at the household level. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

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Figure A7: Country-level UMP

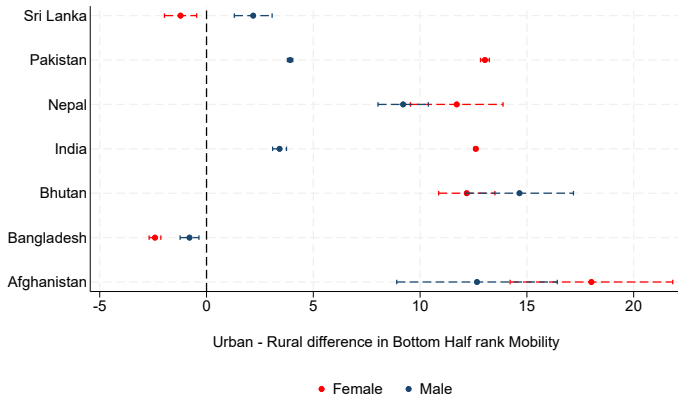
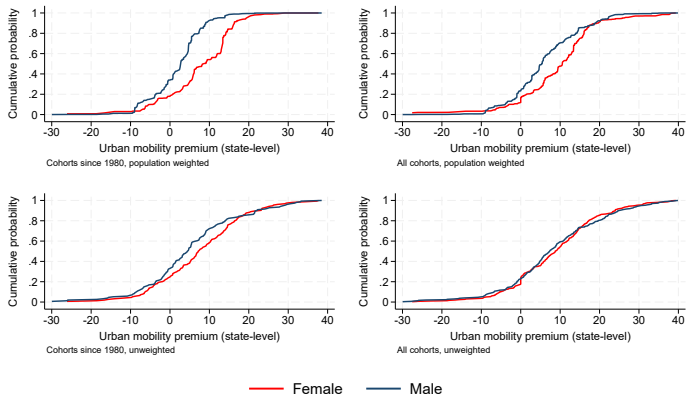


Figure A8: State-level UMP: Robustness to specification



UMP robustness: outcome variable

Figure A9: State-level UMP: Robustness to outcome

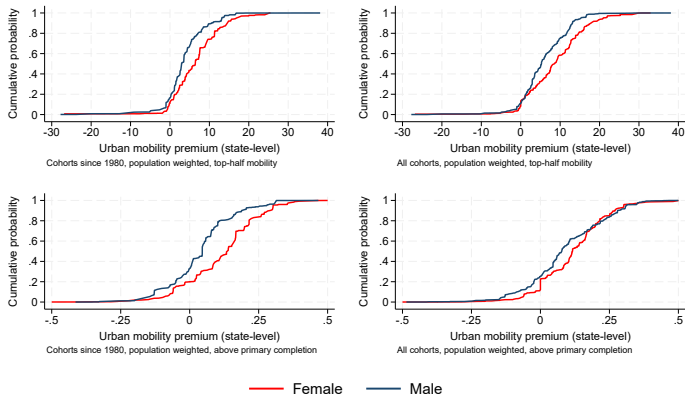


Table A5: Mincer regression: returns to education

| Dependent variable | Log own wage | | | Log husband's wage | |
|---|---------------------|---------------------|---------------------|---------------------|---------------------|
| | (1) | (2) | (3) | (4) | (5) |
| Years of education | 0.080*** (0.000) | 0.069*** (0.000) | 0.051*** (0.000) | 0.036*** (0.000) | 0.019*** (0.001) |
| Years of education \times Female | | 0.016*** (0.000) | 0.007*** (0.001) | | |
| Years of education \times Urban | | | 0.022*** (0.000) | | 0.018*** (0.001) |
| Years of education \times Urban \times Female | | | 0.017*** (0.001) | | |
| Age FE | Yes | Yes | Yes | Yes | Yes |
| Country \times Year FE | Yes | Yes | Yes | Yes | Yes |
| Spouse age | No | No | No | Yes | Yes |
| Observations | 2282207 | 2282207 | 2282207 | 1058529 | 1058529 |
| R^2 | 0.660 | 0.690 | 0.707 | 0.661 | 0.679 |

Note: Robust standard errors in parentheses. Sample in columns (1)-(3) is all working-age (16+) adults in cohorts born since 1980 across six South Asian countries (no data for Bhutan), while sample for columns (4)-(5) is only married women. Wages measured as log daily wage in local currency units. All estimates adjusted with cross-country sampling weights. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

$$\log(w_i) = \alpha + \beta_1 s_i + \beta_2 s_i \times \text{urb}_i + \beta_3 s_i \times \text{fem}_i + \beta_4 s_i \times \text{urb}_i \times \text{fem}_i + \beta_5 \text{urb}_i + \beta_6 \text{fem}_i + \beta_7 \text{fem}_i \times \text{urb}_i + a_a + \delta_{sy} + \epsilon_i$$

Table A6: Bottom-half mobility by returns to education and social norms

| Dependent variable Sample | Child's education rank | | | | | |
|-------------------------------------|------------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
| | Male | | | Female | | |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Return to education | 1.097** (0.445) | | | 1.076*** (0.293) | | |
| Marriage market return to education | | 0.331 (0.552) | | | 1.827* (0.987) | |
| Gender-equal norms index | | | 0.731 (0.894) | | | 3.173*** (1.042) |
| Constant | 34.960*** (1.829) | 38.261*** (1.429) | 38.174*** (1.150) | 34.025*** (2.282) | 35.666*** (2.747) | 36.312*** (1.752) |
| Observations | 126602 | 126602 | 119800 | 93667 | 93667 | 89591 |
| R ² | 0.004 | 0.000 | 0.001 | 0.027 | 0.010 | 0.011 |

Note: Standard errors in parentheses clustered at the state-level. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

[Back-returns](#)
[Back-norms](#)

Table A7: Bottom-half mobility by returns to education and social norms

| Dependent variable Sample | Child's education rank | | | | | |
|------------------------------|------------------------|---------------------|------------------|----------------------|---------------------|---------------------|
| | Male | | | Female | | |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Urban | 3.525*** (1.061) | 0.920 (0.996) | 1.900 (1.564) | 10.536*** (1.608) | 7.360*** (1.653) | 2.771 (2.326) |
| Expenditure rank | | 0.214*** (0.024) | | | 0.236*** (0.020) | |
| Return to education | | | 0.790 (0.707) | | | 1.302*** (0.284) |
| Gender-equal norms index | | | 0.006 (1.053) | | | 3.936*** (1.365) |
| Share of UMP explained | | | 0.461 | | | 0.737 |
| Observations | 112088 | 112088 | 112088 | 82215 | 82215 | 82215 |
| R^2 | 0.003 | 0.050 | 0.005 | 0.030 | 0.086 | 0.053 |

Note: Standard errors in parentheses clustered at the state-level. *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.