



# Outline

## Motivation

- Introduction and Previous work

- Basics

- Examples

## Approach

- Model

- Characterisation

## Inequality Measures

- Main properties

- Example

- Reference point and sensitivity

## Empirical aspects

- Implementation

- Performance

- Application

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# Introduction

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- Many applications proceed just as though cardinal:
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  - but these have limitations
- Present approach based on Cowell and Flachaire (2014)



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# Income Inequality

- 3 ingredients:
  - **“income”**: family income, earnings, wealth  $x \in X \subseteq \mathbb{R}$ .
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- Obviously can't do that here:  $\mu$  is undefined



# Utility

## Cardinalisation and inequality

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  - Already pointed out in Atkinson (1970)
  - Dalton (1920) suggested inequality of (cardinal) utility
  - But if, for all  $i$ , you multiply  $u_i$  by  $\lambda \in (0, 1)$  and add  $\delta = \mu[1 - \lambda]$ ...
  - ...this will automatically reduce measured inequality.



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  - ...this will automatically reduce measured inequality.
- Is this just a technicality?
- Can we proceed just as with regular income?



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# Categorical variable

Example: Access to Services

	Case 1	Case 2
	$n_k$	$n_k$
<u>B</u> oth Gas and Electricity	25	0
<u>E</u> lectricity only	25	50
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- Suppose we have no information about needs / usage
- It seems clear that Case 1 is more unequal than Case 2



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- Compare distributions across countries

## SRH Results: four countries

	<b>Austria</b>	<b>UK</b>	<b>Mexico</b>	<b>Bangladesh</b>
	number of responses			
<i>Very good</i>	423	318	7193	494
<i>Good</i>	390	498	18112	1949
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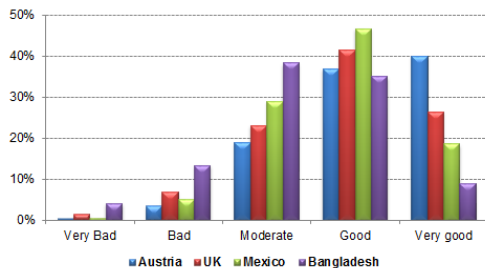
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- For all countries: rank categories in order
- For each country: compute freq distributions across categories
- How to evaluate inequality?

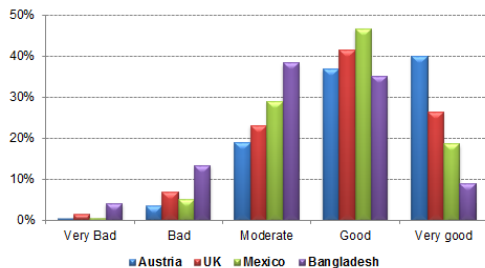
## SRH Inequality: Gini



(1,2,3,4,5)      **At**      **UK**      **Mx**      **BD**      (BD,UK,Mx,At)

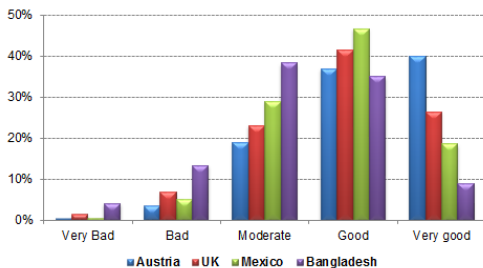
0.111      0.130      0.116      0.154

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(1,2,3,4,5)	0.111	0.130	0.116	0.154	(BD,UK,Mx,At)
(1,2,3,4,1000)	0.593	0.725	0.800	0.884	(BD,Mx,UK,At)

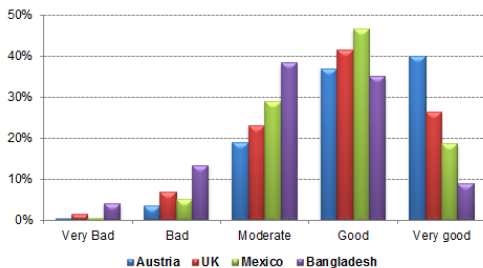
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(1,2,3,4,1000)	0.593	0.725	0.800	0.884	(BD,Mx,UK,At)
(-1000,2,3,4,5)	0.608	0.821	0.856	2.377	(BD,Mx,UK,At)



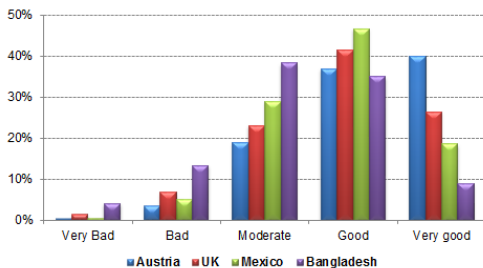
## SRH Inequality: Coeff of Variation



	<b>At</b>	<b>UK</b>	<b>Mx</b>	<b>BD</b>	
(1,2,3,4,5)	0.209	0.244	0.219	0.287	(BD,UK,Mx,At)
(1,2,3,4,1000)	1.210	1.638	2.056	3.088	(BD,Mx,UK,At)



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(1,2,3,4,5)	0.209	0.244	0.219	0.287	(BD,UK,Mx,At)
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(-1000,2,3,4,5)	187.5	11.43	40.45	5.264	(At,Mx,UK,BD)



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# Status and Information



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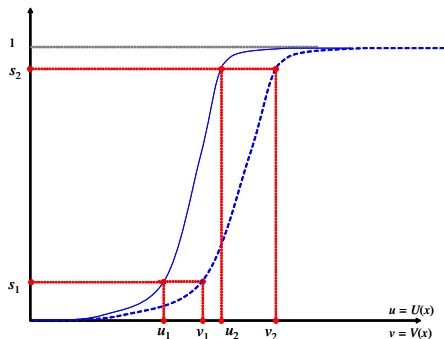
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- In some cases defined given additional distribution-free information
  - example: if it is known that utility is  $\log(x)$
- In some cases requires information on distribution
  - GRE, TOEFL
  - “opportunity” (de Barros et al. 2008)



## Status and Distribution (1)

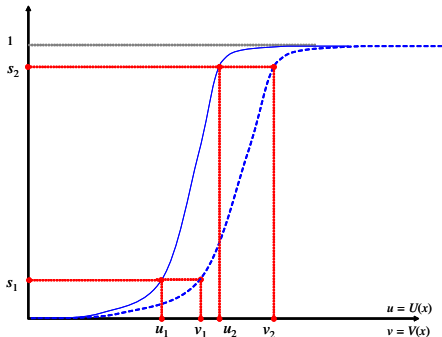
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- disposes of the problem of cardinalisation
  - $U$  and  $V = \varphi(U)$  two cardinalisations of the utility of  $x$
  - for each  $i:u_i$  and  $v_i$  map into  $s_i$



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- This approach works for categorical data
  - we just have an ordered arrangement of categories  $1, 2, \dots, k, \dots, K$
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  - merge two adjacent categories that are irrelevant for  $i$
  - then this should leave  $i$ 's status unaltered
- Principle implies that status should be additive in the  $n_k$ 
  - downward-looking status:  $\sum_{\ell=1}^{k(i)} n_{\ell}$
  - upward-looking status:  $\sum_{\ell=k(i)}^K n_{\ell}$
  - see also Yitzhaki (1979)





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  - $\eta$  need not be increasing in each component of  $\mathbf{s}$
- Inequality: aggregate distance from  $e$ 
  - don't need an explicit distance function
  - implicitly define through inequality ordering  $\succeq$

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- **[Independence]** If  $\mathbf{s}(\zeta, i), \mathbf{s}'(\zeta, i) \in S^n$  satisfy  $(\mathbf{s}(\zeta, i), e) \sim (\mathbf{s}'(\zeta, i), e)$  for some  $\zeta$  then  $(\mathbf{s}(\zeta, i), e) \sim (\mathbf{s}'(\zeta, i), e)$  for all  $\zeta$



## Standard result

### Theorem

*Continuity, Monotonicity, Independence, Anonymity jointly imply  $\succeq$  is representable by the continuous function  $I : S_e^n \rightarrow \mathbb{R}$  where  $I(\mathbf{s}; e) = \Phi(\sum_{i=1}^n d(s_i, e), e)$ , where  $d : S \rightarrow \mathbb{R}$  is a continuous function that is strictly increasing (decreasing) in its first argument if  $s_i > e$  ( $s_i < e$ ).*

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### Corollary

*Inequality is total “distance” from equality. Distance  $d$  is continuous.  $d(s, e)$  is increasing in status if you move away from the reference point.*



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- **[Scale invariance 2]** For all  $\lambda \in \mathbb{R}_+$ : if  $\mathbf{s}, \mathbf{s}', \lambda \mathbf{s}, \lambda \mathbf{s}' \in S^n$  and  $e, e', \lambda e, \lambda e' \in S$  then  $(\mathbf{s}, e) \sim (\mathbf{s}', e') \Rightarrow (\lambda \mathbf{s}, \lambda e) \sim (\lambda \mathbf{s}', \lambda e')$



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### Corollary

Inequality represented as  $I_\alpha(\mathbf{s}; e) := \frac{1}{\alpha[\alpha-1]} \left[ \frac{1}{n} \sum_{i=1}^n s_i^\alpha - e^\alpha \right]$



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## A usable inequality index?

- A *class* of functions available as inequality measures:
  - $\Phi(I_\alpha(\mathbf{s}; e), e)$
  - $e = \eta(\mathbf{s})$ , the reference point
  - $I_\alpha(\mathbf{s}; e) := \frac{1}{\alpha[\alpha-1]} \left[ \frac{1}{n} \sum_{i=1}^n s_i^\alpha - e^\alpha \right]$



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  - $\Phi(I_\alpha(\mathbf{s}; e), e)$
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  - $I_\alpha(\mathbf{s}; e) := \frac{1}{\alpha[\alpha-1]} \left[ \frac{1}{n} \sum_{i=1}^n s_i^\alpha - e^\alpha \right]$
- Do functions  $\Phi(I_\alpha(\mathbf{s}; e), e)$  “look like” inequality measures?
  - transfer principle?
  - reference point?
  - sensitivity to parameters
- What is the appropriate form for  $\Phi$ ?
  - may depend on the reference status  $e$
  - may depend on interpretation

# Outline

## Motivation

- Introduction and Previous work

- Basics

- Examples

## Approach

- Model

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## Inequality Measures

- Main properties

- Example**

- Reference point and sensitivity

## Empirical aspects

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## Summary



## Four distributional scenarios (1)

	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
<b>B</b>	0		25	1	0		25	1
<b>E</b>	50	1	25	3/4	50	1	25	3/4
<b>G</b>	25	1/2	25	1/2	50	1/2	50	1/2
<b>N</b>	25	1/4	25	1/4	0		0	
$\mu(\mathbf{s})$		11/16		5/8		3/4		11/16



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<b>G</b>	25	1/2	25	1/2	50	1/2	50	1/2
<b>N</b>	25	1/4	25	1/4	0		0	
$\mu(\mathbf{s})$		11/16		5/8		3/4		11/16

- $n_k$  is # persons in category  $k \in \{\mathbf{B}, \mathbf{E}, \mathbf{G}, \mathbf{N}\}$



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- $n_k$  is # persons in category  $k \in \{\mathbf{B}, \mathbf{E}, \mathbf{G}, \mathbf{N}\}$
- $s_i = \frac{1}{n} \sum_{\ell=1}^{k(i)} n_{\ell}$  – *downward-looking status*





## Four distributional scenarios (1')

	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s'_i$	$n_k$	$s'_i$	$n_k$	$s'_i$	$n_k$	$s'_i$
<b>B</b>	0		25	1/4	0		25	1/4
<b>E</b>	50	1/2	25	1/2	50	1/2	25	1/2
<b>G</b>	25	3/4	25	3/4	50	1	50	1
<b>N</b>	25	1	25	1	0		0	
$\mu(\mathbf{s})$		11/16		5/8		3/4		11/16



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	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s'_i$	$n_k$	$s'_i$	$n_k$	$s'_i$	$n_k$	$s'_i$
<b>B</b>	0		25	1/4	0		25	1/4
<b>E</b>	50	1/2	25	1/2	50	1/2	25	1/2
<b>G</b>	25	3/4	25	3/4	50	1	50	1
<b>N</b>	25	1	25	1	0		0	
$\mu(\mathbf{s})$		11/16		5/8		3/4		11/16

- $n_k$  is # persons in category  $k \in \{\mathbf{B}, \mathbf{E}, \mathbf{G}, \mathbf{N}\}$



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	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s'_i$	$n_k$	$s'_i$	$n_k$	$s'_i$	$n_k$	$s'_i$
<b>B</b>	0		25	1/4	0		25	1/4
<b>E</b>	50	1/2	25	1/2	50	1/2	25	1/2
<b>G</b>	25	3/4	25	3/4	50	1	50	1
<b>N</b>	25	1	25	1	0		0	
$\mu(\mathbf{s})$		11/16		5/8		3/4		11/16

- $n_k$  is # persons in category  $k \in \{\mathbf{B}, \mathbf{E}, \mathbf{G}, \mathbf{N}\}$
- $s'_i = \frac{1}{n} \sum_{\ell=k(i)}^K n_\ell$  - upward-looking status



## Four distributional scenarios (2)

	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
<b>B</b>	0		25	1	0		25	1
<b>E</b>	50	1	25	3/4	50	1	25	3/4
<b>G</b>	25	1/2	25	1/2	50	1/2	50	1/2
<b>N</b>	25	1/4	25	1/4	0		0	
$\mu(\mathbf{s})$		11/16		5/8		3/4		11/16

- Case 0 to Case 1:



## Four distributional scenarios (2)

	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
<b>B</b>	0		25	1	0		25	1
<b>E</b>	50	1	25	3/4	50	1	25	3/4
<b>G</b>	25	1/2	25	1/2	50	1/2	50	1/2
<b>N</b>	25	1/4	25	1/4	0		0	
$\mu(s)$	11/16		5/8		3/4		11/16	

- Case 0 to Case 1:
  - 25 people promoted from E to B
  - if  $e$  equals to any of values taken by  $\mu(s)$
  - then inequality increases



## Four distributional scenarios (3)

	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
<b>B</b>	0		25	1	0		25	1
<b>E</b>	50	1	25	$3/4$	50	1	25	$3/4$
<b>G</b>	25	$1/2$	25	$1/2$	50	$1/2$	50	$1/2$
<b>N</b>	25	$1/4$	25	$1/4$	0		0	
$\mu(\mathbf{s})$		$11/16$		$5/8$		$3/4$		$11/16$

- Case 0 to Case 2:



## Four distributional scenarios (3)

	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
<b>B</b>	0		25	1	0		25	1
<b>E</b>	50	1	25	3/4	50	1	25	3/4
<b>G</b>	25	1/2	25	1/2	50	1/2	50	1/2
<b>N</b>	25	1/4	25	1/4	0		0	
$\mu(s)$	11/16		5/8		3/4		11/16	

- Case 0 to Case 2:
  - 25 people promoted from N to G
  - if  $e$  equals to any of values taken by  $\mu(s)$
  - then inequality decreases



# “Transfer Principle”?





## “Transfer Principle”?

	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
<b>B</b>	0		25	1	0		25	1
<b>E</b>	50	1	25	3/4	50	1	25	3/4
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	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
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- Case 0 to Case 1: inequality increases



## “Transfer Principle”?

	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
<b>B</b>	0		25	1	0		25	1
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<b>N</b>	25	1/4	25	1/4	0		0	
$\mu(\mathbf{s})$		11/16		5/8		3/4		11/16

- Case 0 to Case 1: inequality increases
- Case 0 to Case 2: inequality decreases



## “Transfer Principle”?

	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
<b>B</b>	0		25	1	0		25	1
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<b>N</b>	25	1/4	25	1/4	0		0	
$\mu(\mathbf{s})$		11/16		5/8		3/4		11/16

- Case 0 to Case 1: inequality increases
- Case 0 to Case 2: inequality decreases
- Case 0 to Case 3: combination results in ambiguous change



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## Reference point



## Reference point

- **Mean status:**  $e = \eta(\mathbf{s}) = \mu(\mathbf{s})$ 
  - for continuous distributions will equal 0.5
  - for categorical data, there is no counterpart to fixed-mean assumption in income-inequality analysis



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- **Mean status:**  $e = \eta(\mathbf{s}) = \mu(\mathbf{s})$ 
  - for continuous distributions will equal 0.5
  - for categorical data, there is no counterpart to fixed-mean assumption in income-inequality analysis
- **Median status:**  $e = \eta(\mathbf{s}) = \text{med}(\mathbf{s})$ 
  - not well-defined: any value in interval  $M(\mathbf{s})$
  - $M(\mathbf{s}) = [1/2, 1)$  in cases 0 and 2
  - $M(\mathbf{s}) = [1/2, 3/4)$  in cases 1 and 3





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- **Max status:**  $e = 1$ 
  - for constant  $e$  this is only value that makes sense



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  - $M(\mathbf{s}) = [1/2, 3/4)$  in cases 1 and 3
- **Max status:**  $e = 1$ 
  - for constant  $e$  this is only value that makes sense
- **Min status:**  $e = 0$ 
  - counterpart for peer-exclusive case

# Sensitivity

- $\alpha$  captures the sensitivity of measured inequality



## Sensitivity

- $\alpha$  captures the sensitivity of measured inequality
- If  $\alpha$  is high  $I_\alpha(\mathbf{s}; e) = \frac{1}{\alpha[\alpha-1]} \left[ \frac{1}{n} \sum_{i=1}^n s_i^\alpha - e^\alpha \right]$ , sensitive to high status-inequality



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- If  $\alpha = 0$  then  $I_0(\mathbf{s}; e) = -\frac{1}{n} \sum_{i=1}^n \log s_i + \log e$ ,



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- If  $\alpha = 0$  then  $I_0(\mathbf{s}; e) = -\frac{1}{n} \sum_{i=1}^n \log s_i + \log e$ ,
- If  $e = \mu(\mathbf{s})$  and  $\alpha = 1$  then  $\frac{1}{n} \sum_{i=1}^n s_i \log s_i - e \log e$

## Behaviour of $I_0(\mathbf{s}; e)$

	Case 0	Case 1	Case 2	Case 3
$\mu(\mathbf{s})$	11/16	5/8	3/4	11/16
$\text{med}_1(\mathbf{s})$	3/4	5/8	3/4	5/8
$\text{med}_2(\mathbf{s})$	1/2	1/2	1/2	1/2
$I_0(\mathbf{s}; \mu(\mathbf{s}))$	0.1451	0.1217	0.0588	0.0438
$I_0(\mathbf{s}; \text{med}_1(\mathbf{s}))$	0.2321	0.1217	0.0588	-0.0515
$I_0(\mathbf{s}; \text{med}_2(\mathbf{s}))$	-0.1732	-0.1013	-0.3465	-0.2746
$I_0(\mathbf{s}; 1)$	0.5198	0.5917	0.3465	0.4184



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- $I_0(\mathbf{s}; \mu(\mathbf{s})), I_0(\mathbf{s}; \text{med}_1(\mathbf{s}))$ : inequality *decreases* for
  - Case 0 to 1, or Case 2 to 3
  - movement changes both the  $\mu(\mathbf{s})$  and  $\text{med}_1(\mathbf{s})$  ref points





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- $I_0(\mathbf{s}; \text{med}_2(\mathbf{s})) < 0$  for *all* cases in example!
- But  $I_0(\mathbf{s}; 1)$  seems sensible



## Inequality measure

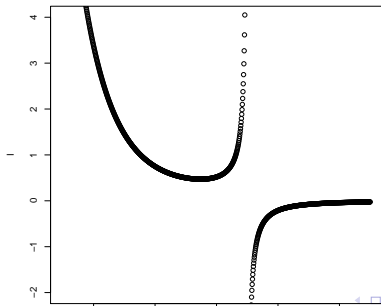
- For ordinal data, peer-inclusive status



## Inequality measure

- For ordinal data, peer-inclusive status

$$I_{\alpha}(\mathbf{s}, 1) = \begin{cases} \frac{1}{\alpha(\alpha-1)} \left[ \frac{1}{n} \sum_{i=1}^n s_i^{\alpha} - 1 \right], & \text{if } \alpha \neq 0, \alpha < 1 \\ -\frac{1}{n} \sum_{i=1}^n \log s_i. & \text{if } \alpha = 0 \end{cases}$$



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# Implementation

- Description of sample

$$x_i = \begin{cases} 1 & \text{with sample proportion } p_1 \\ 2 & \text{with sample proportion } p_2 \\ \dots & \\ K & \text{with sample proportion } p_K \end{cases},$$



## Implementation

- Description of sample

$$x_i = \begin{cases} 1 & \text{with sample proportion } p_1 \\ 2 & \text{with sample proportion } p_2 \\ \dots & \\ K & \text{with sample proportion } p_K \end{cases},$$

- Point estimate of the index:

$$I_\alpha = \begin{cases} \frac{1}{\alpha(\alpha-1)} \left[ \sum_{i=1}^K p_i \left[ \sum_{j=1}^i p_j \right]^\alpha - 1 \right] & \text{if } \alpha \neq 0,1 \\ - \sum_{i=1}^K p_i \log \left[ \sum_{j=1}^i p_j \right] & \text{if } \alpha=0 \end{cases}$$

- function of  $K$  parameter estimates  $(p_1, p_2, \dots, p_K)$  following a multinomial



# Asymptotics





# Asymptotics

- From the CLT  $I_\alpha$  is asymptotically Normally distributed

## Asymptotics

- From the CLT  $I_\alpha$  is asymptotically Normally distributed
- Estimator of cov matrix of  $(p_1, p_2, \dots, p_k)$  is

$$\Sigma = \frac{1}{n} \begin{bmatrix} p_1(1-p_1) & -p_1p_2 & \dots & -p_1p_K \\ -p_2p_1 & p_2(1-p_2) & \dots & -p_2p_K \\ \vdots & \vdots & \vdots & \vdots \\ -p_Kp_1 & -p_Kp_2 & \dots & p_K(1-p_K) \end{bmatrix}$$





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# Confidence Intervals

## Confidence Intervals

- 3 variants of CIs: Asymptotic, Percentile Bootstrap, Studentized Bootstrap
- $CI_{asym} = [I_\alpha - c_{0.975} \widehat{\text{Var}}(I_\alpha)^{1/2}; I_\alpha + c_{0.975} \widehat{\text{Var}}(I_\alpha)^{1/2}]$ 
  - $c_{0.975}$  from the Student distribution  $T(n-1)$
  - do not always perform well in finite samples



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- Bootstraps: generate resamples,  $b = 1, \dots, B$ 
  - for each resample  $b$  compute the inequality index
  - obtain  $B$  bootstrap statistics,  $I_\alpha^b$
  - also  $B$  bootstrap  $t$ -statistics  $t_\alpha^b = (I_\alpha^b - I_\alpha) / \widehat{\text{Var}}(I_\alpha^b)^{1/2}$



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  - also  $B$  bootstrap  $t$ -statistics  $t_\alpha^b = (I_\alpha^b - I_\alpha) / \widehat{\text{Var}}(I_\alpha^b)^{1/2}$
- $CI_{perc} = [c_{0.025}^b; c_{0.975}^b]$ 
  - $c_{0.025}^b$  and  $c_{0.975}^b$  are from EDF of bootstrap statistics





## Confidence Intervals

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  - do not always perform well in finite samples
- Bootstraps: generate resamples,  $b = 1, \dots, B$ 
  - for each resample  $b$  compute the inequality index
  - obtain  $B$  bootstrap statistics,  $I_\alpha^b$
  - also  $B$  bootstrap  $t$ -statistics  $t_\alpha^b = (I_\alpha^b - I_\alpha) / \widehat{\text{Var}}(I_\alpha^b)^{1/2}$
- $CI_{perc} = [c_{0.025}^b; c_{0.975}^b]$ 
  - $c_{0.025}^b$  and  $c_{0.975}^b$  are from EDF of bootstrap statistics
- $CI_{stud} = [I_\alpha - c_{0.975}^* \widehat{\text{Var}}(I_\alpha)^{1/2}; I_\alpha - c_{0.025}^* \widehat{\text{Var}}(I_\alpha)^{1/2}]$ 
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# Performance Test



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- Check coverage error rate of CIs as sample size increases
  - $\alpha = -1, 0, 0.5, 0.99$
  - 199 bootstraps
  - 10 000 replications to compute error rates
  - $n = 20, 50, 100, 200, 500, 1000$

## Estimation Methods Compared

	$\alpha$	-1	0	0.5	0.99
Asymptotic B	$n = 20$	0.0606	0.0417	0.0598	0.0491
	$n = 500$	0.0523	0.0492	0.0521	0.0523
	$n = 1000$	0.0485	0.0540	0.0552	0.0549
Percentile B	$n = 20$	0.0384	0.0981	0.0912	0.1023
	$n = 500$	0.0509	0.0513	0.0552	0.0554
	$n = 1000$	0.0482	0.0556	0.0547	0.0551
Studentized B	$n = 20$	0.1275	0.0843	0.1041	0.1377
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- Asymptotic CIs perform OK in finite sample
- Percentile bootstrap performs well for  $n > 50$
- Studentized bootstrap does not do well for small samples
- Reliable results for  $\alpha = 0.99$  (index is undefined for  $\alpha = 1$ .)

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Examples

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Model

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## Inequality Measures

Main properties

Example

Reference point and sensitivity

## Empirical aspects

Implementation

Performance

**Application**

## Summary



# World Values Survey

## World Values Survey

- Life satisfaction question:

*All things considered, how satisfied are you with your life as a whole these days? Using this card on which 1 means you are “completely dissatisfied” and 10 means you are “completely satisfied” where would you put your satisfaction with your life as a whole? (code one number):*

*Completely dissatisfied – 1 2 3 4 5 6 7 8 9 10 – Completely satisfied*

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- Health question:

*All in all, how would you describe your state of health these days? Would you say it is (read out):*

*1 Very good, 2 Good, 3 Fair, 4 Poor.*

## GDP and Life satisfaction

- Cross-country comparison of life satisfaction and GDP/head
  - happiness-income paradox (Easterlin 1974, Clark and Senik 2011)
  - weak relation happiness-income internationally? (Easterlin 1995, Easterlin et al. 2010)
  - or a strong relationship? (Hagerty and Veenhoven 2003, Deaton 2008, Stevenson and Wolfers 2008a, Inglehart et al. 2008)

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- How should we quantify life satisfaction?
  - simple linearity of Likert scale? or exponential scale?
  - Ng (1997), Ferrer-i-Carbonell and Frijters (2004), Kristoffersen (2011)



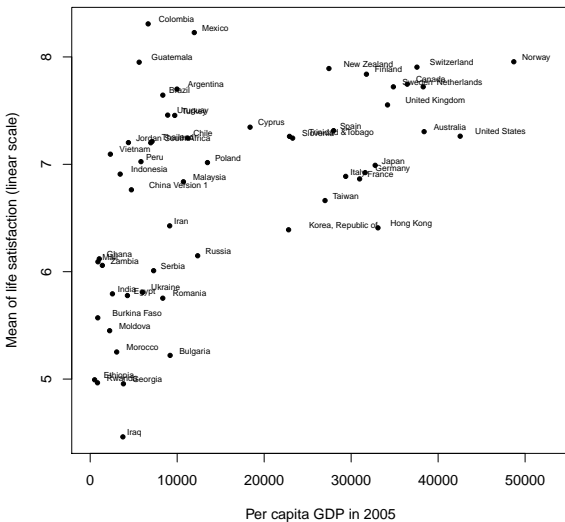


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- Is inequality of life satisfaction related to GDP/head?
  - Use  $I_0$  and other members of the same family

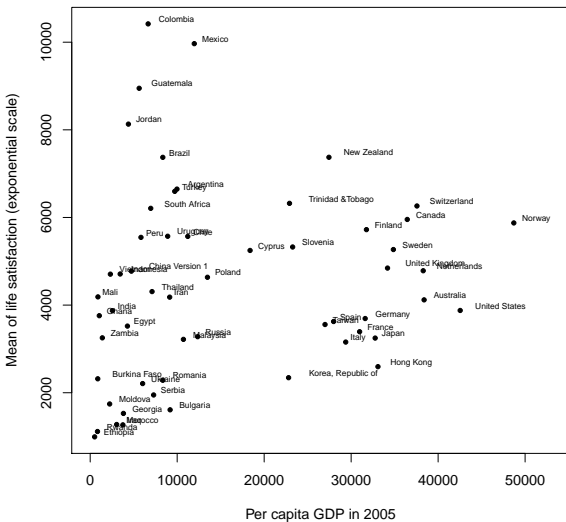


# GDP and Life satisfaction (Linear)

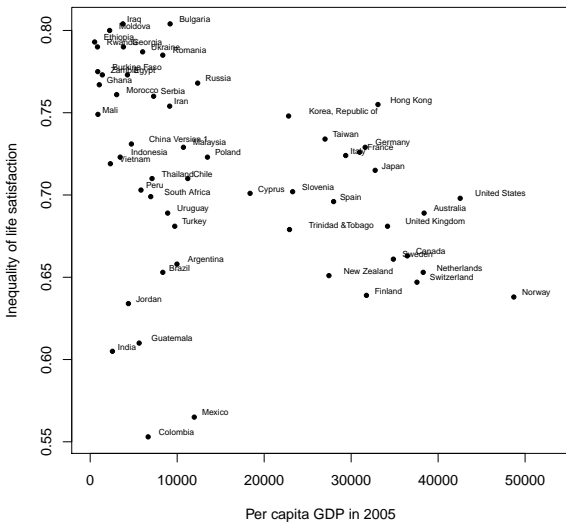




# GDP and Life satisfaction (Exponential)

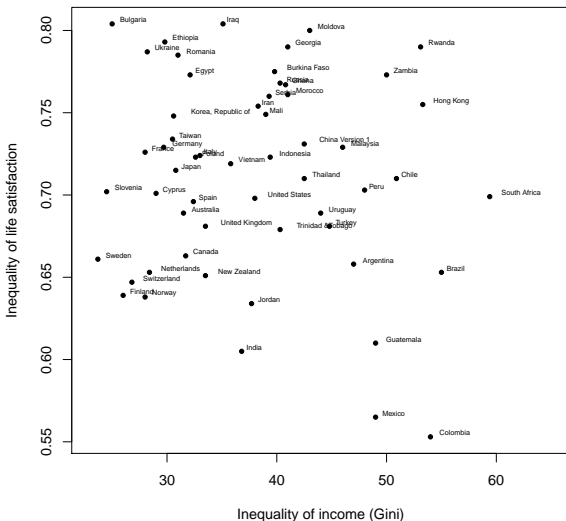


# GDP and Inequality of Life satisfaction





# Income inequality and Inequality of Life satisfaction





## Health status

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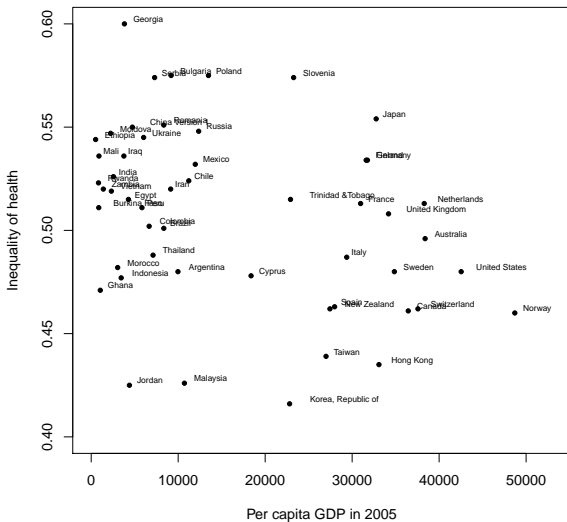
## Health status

- Health is HRS
- Cross-country comparison of health and GDP
  - a significant positive relationship? (Deaton 2008)
- Cross-country comparison of inequality of health and Inequality of life satisfaction
  - use same inequality index as for life satisfaction



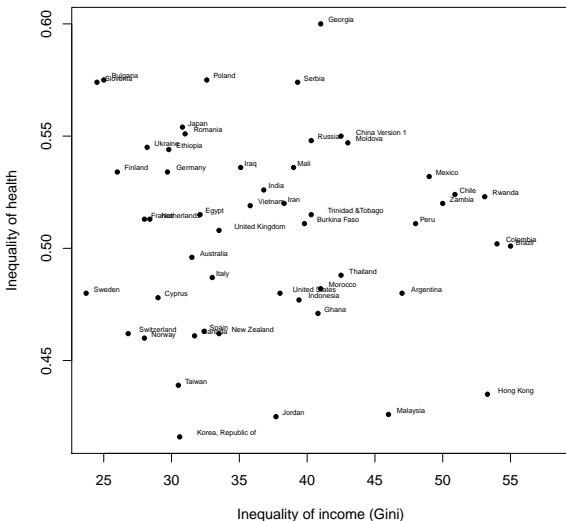


# GDP and Inequality of health



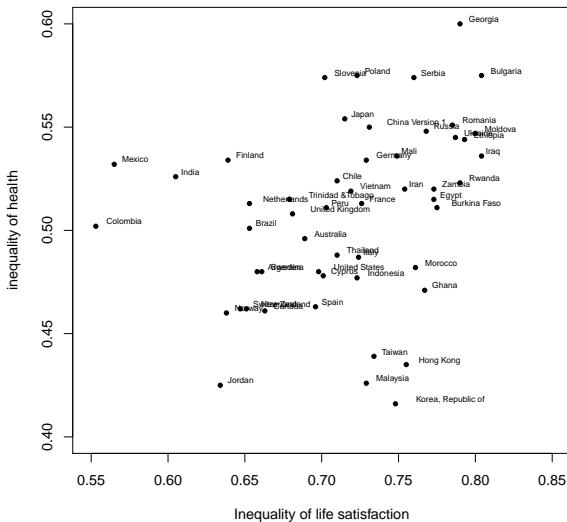


# Income inequality and health inequality





# Inequality of life satisfaction and health inequality





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- No clear relationship between  $I_0$ (health) on GDP per capita
- OLS estimate of  $I_0$ (health) on  $I_0$ (life satisfaction) produces a slope coefficient not significantly different from zero
  - health-life satisfaction relationship is not significant





## Summary

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  - separates out the issue of status from that of inequality-aggregation
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  - gives a family of measures
- Nice properties empirically



# Summary\_

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## Median: definition and in our cases

- $\text{med}(\mathbf{s})$  defined as  $e \in S$  such that  $\#(s_i \leq e) \geq \frac{n}{2}$ ,  $\#(s_i \geq e) \geq \frac{n}{2}$



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	Case 0		Case 1		Case 2		Case 3	
	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$	$n_k$	$s_i$
<b>B</b>	0		25	1	0		25	1
<b>E</b>	50	1	25	3/4	50	1	25	3/4
<b>G</b>	25	1/2	25	1/2	50	1/2	50	1/2
<b>N</b>	25	1/4	25	1/4	0		0	
$M(\mathbf{s})$	[1/2, 1)		[1/2, 3/4)		[1/2, 1)		[1/2, 3/4)	

- $\text{med}(\mathbf{s})$  could be any value in interval  $M(\mathbf{s})$





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- Three ordered categories
- Same proportion of individuals in each category



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  - $2/3$  of the population has a status greater than or equal to  $m$
- Median as “half-way” point is misleading



## Median example 2

- Two ordered categories ( B better than A)
- Three distributions
  1.  $n_A = 500, n_B = 500$
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- Compare:
  - distributions 1 and 2 have very different medians
  - distributions 2 and 3 have almost the same median!

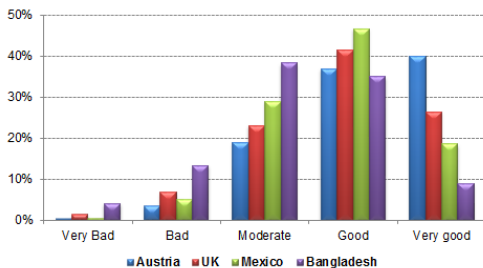


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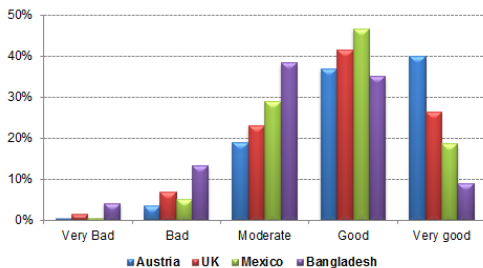
## SRH Inequality: Gini (median norm'd)



(1,2,3,4,5)      **At**      **UK**      **Mx**      **BD**      (BD,UK,Mx,At)\*

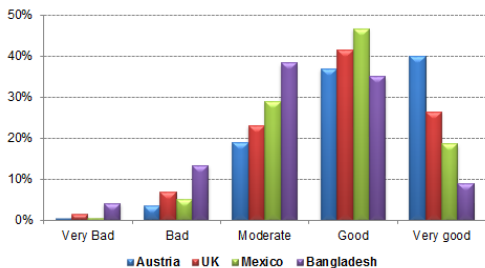
0.107      0.135      0.123      0.140

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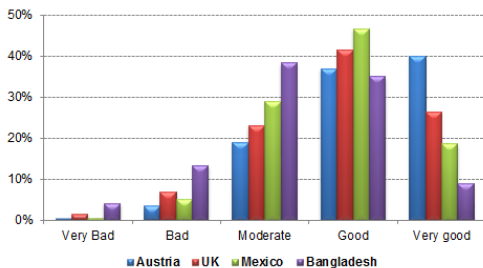
	<b>At</b>	<b>UK</b>	<b>Mx</b>	<b>BD</b>	
(1,2,3,4,5)	0.107	0.135	0.123	0.140	(BD,UK,Mx,At)*
(1,2,3,4,1000)	0.006	0.011	0.017	0.029	(BD,Mx,UK,At)*

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(1,2,3,4,1000)	0.006	0.011	0.017	0.029	(BD,Mx,UK,At)*
(-1000,2,3,4,5)	7.39	-0.315	-1.844	-0.188	(At,Mx,UK,BD)

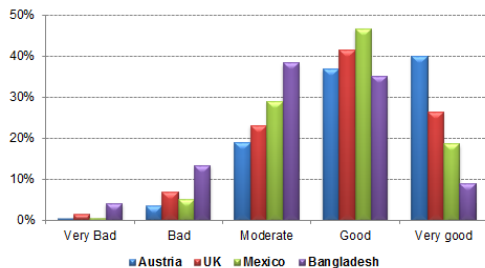
## SRH Inequality: C of V (median norm'd)



(1,2,3,4,5)      **At**      **UK**      **Mx**      **BD**      (BD,UK,Mx,At) \*

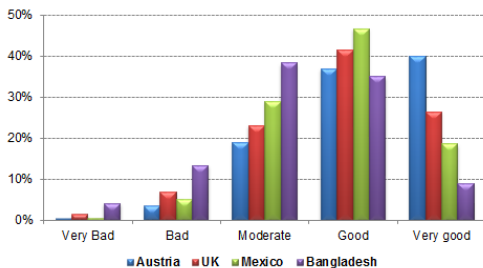
0.202      0.253      0.232      0.260

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	<b>At</b>	<b>UK</b>	<b>Mx</b>	<b>BD</b>	
(1,2,3,4,5)	0.202	0.253	0.232	0.260	(BD,UK,Mx,At) *
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(1,2,3,4,1000)	0.012	0.024	0.044	0.101	(BD,Mx,UK,At)*
(-1000,2,3,4,5)	2276	-4.39	-87.2	-0.42	(At,BD,UK,Mx)



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## Proof of Theorem 1

- Two cases to consider
  - data are categorical:  $S$  is set of non-negative rational numbers,  $\mathbb{Q}_+$ .
  - data have cardinal significance:  $S$  can be taken as an interval in  $\mathbb{R}$ .
- In either case  $(S, +, \succ)$  forms a strictly ordered group (Krantz 1964, Luce and Tukey 1964, Wakker 1988)
- From Theorem 5.3 of Fishburn (1970) Axioms jointly imply that, for a given  $e$ ,  $\succ$  is representable by a continuous function  $S^{n+1} \rightarrow \mathbb{R}: \sum_{i=1}^n d_i(s_i, e), \forall (\mathbf{s}, e) \in S^{n+1}$  where, for each  $i$ ,  $d_i : S \rightarrow \mathbb{R}$  is a continuous function.
- By monotonicity this is increasing in  $s_i$  if  $s_i > e$  and *vice versa*.
- By anonymity the functions  $d_i$  must all be identical
- ordering  $\succ$  is also representable any monotonic transform





## “Maximum inequality”

- Take the case where status is downward-looking and peer-inclusive
- Suppose that the status of each member of category  $k$  is  $s$
- If a person is promoted from category  $k$  to category  $k + 1$ 
  - status increases to  $s + n_{k+1}/n$
  - status of each of the remaining  $n_k - 1$  members of category  $k$  falls to  $s - 1/n$ .
- The resulting change in inequality is proportional to
 
$$\left[ d\left(s + \frac{n_{k+1}}{n}, e\right) - d(s, e) \right] + [n_k - 1] \left[ d\left(s - \frac{1}{n}, e\right) - d(s, e) \right]$$
- If  $d$  is differentiable then this expression is approximately
 
$$d'(s, e) \frac{n_{k+1}}{n} - \frac{n_k - 1}{n} d'(s, e)$$
  - which equals  $\frac{1}{n} d'(s, e) [n_{k+1} - n_k + 1]$ .
- If  $s < e$  then monotonicity implies  $d'(s, e) < 0$ 
  - the change in inequality is negative if  $n_{k+1} \geq n_k$ .

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## Dispersion

- Model:  $\text{LifeSatisf}_i = \alpha + \beta \text{GDP}_i + \varepsilon_i$ ,  $\varepsilon_i \sim N(0, \sigma_i^2)$ 
  - $\beta$  is a significant coefficient and  $R^2$  is large
  - strong (linear) relationship between LifeSatisf and GDP
- If LifeSatisf equation is homoskedastic:
  - no relationship between GDP and the dispersion of LifeSatisf
  - whatever is GDP, the dispersion of LifeSatisf is the same
- If LifeSatisf equation heteroskedastic dispersion of LifeSatisf may or may not be related to GDP
  - the form of the heteroskedasticity cannot be deduced from the relationship between the dependent variable and the covariate.
- If every  $i$  has different GDP,  $\sigma_i^2$  measures the dispersion of LifeSatisf for  $i$
- taking the measure  $I_0$  as a measure of dispersion, the same reasoning applies

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